Parental age, genetic mutation, and cerebral palsy

N A Fletcher, J Foley

Abstract
Parental age and birth order were studied in 251 patients with cerebral palsy. No parental age or birth order effects were observed in spastic quadriplegia or diplegia, but a paternal age effect was detected in those with athetoid/dystonic cerebral palsy and congenital hemiplegia. These observations indicate that some cases of athetoid/dystonic or hemiplegic cerebral palsy might arise by fresh dominant genetic mutation.

Methods
All cases were ascertained from the records of the Cheyne Centre for Spastic Children, London. All had been referred from general practitioners or paediatricians in south-east England and were seen and examined personally by one of the authors (JF) between 1955 and 1986. CP was defined as above, low birth weight (LBW) as < 2500 g, and prematurity as birth before 37 weeks' gestation.

Cases with recognisable syndromes or malformations, genetic or metabolic disorders, kernicterus, other identifiable causes for the motor disorder (including postnatal factors), similarly affected relatives, ataxic CP, or asymmetrical quadriplegia were excluded.

The records of 532 cases were examined. Of these, 172 had spastic diplegia (65% preterm, 69% LBW), 102 had congenital hemiplegia (23% preterm, 20% LBW), 104 had symmetrical spastic quadriplegia (26% preterm, 35% LBW), and 154 had athetoid/dystonic CP (14% preterm, 19% LBW). Sibship size, birth order, parental ages at the time of birth, and date of birth were recorded. Of the 532 cases, 399 were born in the UK to parents whose ages had been recorded; 251 of these were born between 1961 and 1983 which is the period for which the Registrar General has published annual mean parental ages for the UK population.

Methods of birth order and parental age analysis were those described by Emery.1 Birth order was analysed using the method of Haldane and Smith.12 For each case, the observed birth order is recorded and the expected birth order and its variance (assuming no birth order effect) calculated from the sibship size and published tables.11 In each group of cases, the standard error of the difference between the sums of the observed and expected birth orders is the square root of the sum of the variances. Parental ages at the time of birth of each case were compared with mean UK population maternal and paternal ages for the corresponding year using paired t tests. For each group of cases, mean differences between observed and expected values of paternal age (d_p), maternal age (d_m), and birth order (d_o) were calculated.

As d_p, d_m, and d_o are mutually dependent variables, a method based on regression analysis was used to compare any statistically significant value of one with that predicted by the observed values of the other two.13 The difference between observed and predicted values is d_o−b_3 d_p, where d_o is the mean increase in one of the three variables (d_p, d_m, or d_o), d_p is the mean increase in one of the other two, and b_3 is a regression coefficient given by s_p/s_o.
where \( s_1 \) and \( s_2 \) are the standard deviations and \( r_{12} \) is the product moment correlation coefficient of variables 1 and 2 in the population. Standard errors for these differences can be calculated as described by Smith.\(^3\) For these calculations, the following correlation coefficients were used: birth order and paternal age = 0.3, birth order and maternal age = 0.45, and maternal age and paternal age = 0.72.\(^11\) Population standard deviations for paternal age (64 years), maternal age (54 years), and birth order (1-31) were assumed to be those in the Registrar General's statistics for 1970.\(^11\)

### Results

Of the 251 cases included, 50 had spastic quadriplegia, 49 hemiplegia, 75 diplegia, and 77 athetoid/dystonic CP. Parental age data for each group of patients are given in Table 1. In those with quadriplegia there was no significant increase in paternal age (\( d_p = +0.99 \) years, \( p > 0.2 \)), maternal age (\( d_m = +0.2 \) years, \( p > 0.7 \)), or birth order (\( d_b = +0.12, p > 0.1 \)). In the diplegia group, paternal age (\( d_p = +1.27 \) years, \( p > 0.05 \)), maternal age (\( d_m = -0.33 \) years), and birth order (\( d_b = +0.18, p > 0.1 \)) were also similar to those in the UK population. No significant paternal age effect was seen in the diplegia cases when preterm (+1.19 years, \( n = 51 \)) and term (+1.45 years, \( n = 24 \)) births were analysed separately.

Significant parental age and birth order effects were seen in those with athetoid and hemiplegic CP and these data are shown in Table 2. In the hemiplegia group (Table 2), birth order was not significantly increased after allowance for the observed increases in parental ages. Paternal age was not significantly increased after adjustment for the observed rise in maternal age and vice versa. However, the paternal age effect was greater and remained (at the 0.05 level) after adjustment for the observed rise in birth order, while this was not so for maternal age.

In the athetoid/dystonic CP cases (Table 2), a birth order effect also disappeared after allowing for either parental age effect. Although neither parental age increase was significant after correction for the other, there was a highly significant difference between the paternal age effect and that predicted by the rise in birth order, while this was barely significant for maternal age.

### Discussion

Until recently, CP was thought to be caused by adverse perinatal factors, especially hypoxic-ischaemic damage.\(^1\) This has been accepted particularly in athetoid/dystonic CP.\(^14\) Although some prenatal and perinatal risk factors are associated with the subsequent development of CP, many of these are not specific to asphyxia,\(^15\) nor necessarily causative, and absent in the majority of children with CP. Moreover, the majority of those in whom such risk factors are present develop normally.\(^1\)

To test our hypothesis that some isolated cases of spastic or athetoid/dystonic CP arise because of genetic mutation, we have tried to determine whether paternal age among cases is higher than in the normal population. A serious difficulty in birth order and parental age studies arises from the interrelationship between the three variables; a change in one is associated with an alteration of one or both of the others. All statistical approaches to this difficulty are open to methodological criticism\(^11\) and it may be difficult or impossible to establish which variable is increased because of an association with the cause of a disorder.\(^1\) In this study, birth order and parental age are significantly increased only among cases with athetoid/dystonic CP and congenital hemiplegia. In athetoid/dystonic CP, maternal and paternal age are still increased after allowing for birth order but the paternal age increase is more significant. Therefore, a paternal age effect seems more likely than a maternal one. In the hemiplegia group, the paternal age effect is just significant after birth order is taken into account. The results indicate a paternal age effect in athetoid/dystonic CP and possibly in hemiplegic CP.

There are other reasons for doubting that increased maternal age is the more important factor. CP is thought to be related to parity or maternal age in a similar way to perinatal mortality.\(^16\) The latter is associated with younger (less than 20 years) as well as older (over 34 years old) mothers and is higher in first born infants than in later birth orders, only rising again with high parities of five or more.\(^17,18\) An association between total CP and increased maternal age was noted in an earlier series,\(^19\) but paternal age was not examined; consequently this could have been secondary to a paternal age effect.

A genetic contribution to the causation of CP has long been suspected; several genetic disorders may resemble CP clinically\(^20\) and

### Table 1 Parental age data.

<table>
<thead>
<tr>
<th>CP subgroup</th>
<th>Mean parental age (y)</th>
<th>SD</th>
<th>Range</th>
</tr>
</thead>
<tbody>
<tr>
<td>Quadruplegia (n = 50)</td>
<td>Paternal</td>
<td>30-2</td>
<td>5-74</td>
</tr>
<tr>
<td></td>
<td>Maternal</td>
<td>27-1</td>
<td>4-92</td>
</tr>
<tr>
<td>Diplegia (n = 75)</td>
<td>Paternal</td>
<td>30-7</td>
<td>6-73</td>
</tr>
<tr>
<td></td>
<td>Maternal</td>
<td>26-8</td>
<td>6-62</td>
</tr>
<tr>
<td>Athetoid (n = 77)</td>
<td>Paternal</td>
<td>31-4</td>
<td>7-33</td>
</tr>
<tr>
<td></td>
<td>Maternal</td>
<td>28-7</td>
<td>6-01</td>
</tr>
<tr>
<td>Hemiplegia (n = 49)</td>
<td>Paternal</td>
<td>31-5</td>
<td>7-12</td>
</tr>
<tr>
<td></td>
<td>Maternal</td>
<td>28-9</td>
<td>6-14</td>
</tr>
</tbody>
</table>

### Table 2 Parental age and birth order analysis in athetoid and hemiplegic CP.

<table>
<thead>
<tr>
<th>CP subgroup</th>
<th>Observed value of ( d_p/d_m/d_b )</th>
<th>Observed value minus that predicted by value of ( d_p )</th>
</tr>
</thead>
<tbody>
<tr>
<td>Athetoid (n = 77)</td>
<td>( d_p = +2.16 ) (p &lt; 0.02)</td>
<td>+0.66 (NS) +1.78 (p &lt; 0.02)</td>
</tr>
<tr>
<td></td>
<td>( d_m = +1.77 ) (p &lt; 0.02)</td>
<td>+0.45 (NS) +1.13 (p &lt; 0.02)</td>
</tr>
<tr>
<td></td>
<td>( d_b = +0.26 ) (p &lt; 0.01)</td>
<td>+0.07 (NS) - (NS)</td>
</tr>
</tbody>
</table>

| Hemiplegia (n = 49) | \( d_p = +2.60 \) (p < 0.02) | +0.90 (NS) +2.17 (p < 0.02) |
| | \( d_m = +0.43 \) (p < 0.05) | +1.48 (NS) +1.48 |
| | \( d_b = +0.33 \) (p < 0.02) | +0.07 (NS) - (NS) |

\( d_p = \) paternal age effect, \( d_m = \) maternal age effect, \( d_b = \) birth order effect, NS = not significant (p > 0.05).
Whether mutations have been reported in athetoid/dystonic CP, and recurrence risks for sibs may be as high as 10% if signs are symmetrical and birth history normal. Whether the normal birth history criterion is any longer valid in the light of recent studies of perinatal factors in CP is uncertain.

The presence of a paternal age effect in athetoid/dystonic CP suggests that some apparently isolated cases may be the result of a new genetic mutation; this may apply to hemiplegic CP but the evidence in this group is weaker. Athetoid/dystonic and hemiplegic CP are clinically and pathologically heterogeneous, so that cases resulting from new genetic mutations may only be a subgroup. The existence of paternally derived CP mutations is an intriguing possibility in the light of the apparently vascular basis of some pathological lesions in athetoid/dystonic CP and hemiplegia. Owing to genomic imprinting effects, paternal genes are critical for the development and maintenance of the placent. Could placental dysfunction as a result of a paternally derived genetic mutation cause fetal cerebrovascular damage leading directly (or in combination with otherwise insufficient intrapartum factors) to athetoid/dystonic CP and possibly congenital hemiplegia? If fresh dominant mutation predisposes to affected sibs and is caused by some cases of athetoid/dystonic CP, it would seem likely and some affected children would be predicted. Careful family studies with unbiased ascertainment, examination of offspring, and estimation of reproductive fitness of CP patients would be a useful means of studying this issue further.

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